

Best linear estimator - The solution

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(a) We have:

$$y_i = \beta x_i + \epsilon_i$$

Let $\hat{\beta} = \sum_i a_i y_i$ be any linear estimator.

So,

$$\mathbb{E}(\hat{\beta}) = \sum_i a_i \mathbb{E}(y_i) = \beta \sum_i a_i x_i$$

and,

$$\text{Bias}(\hat{\beta}) = \mathbb{E}(\hat{\beta}) - \beta = \beta \left[\sum_i a_i x_i - 1 \right].$$

Similarly,

$$\text{var}(\hat{\beta}) = \sum_i a_i^2 \text{var}(y_i) = \sigma^2 \sum_i a_i^2.$$

So,

$$\begin{aligned} \text{MSE}(\hat{\beta}) &= M = \sigma^2 \sum_i a_i^2 + \beta^2 \left[\sum_i a_i x_i - 1 \right]^2. \\ \frac{\partial M}{\partial a_j} &= 2\sigma^2 a_j + 2\beta^2 \left[\sum_i a_i x_i - 1 \right] x_j = 0, \quad \forall j \end{aligned} \tag{1}$$

Multiply by y_j and add over all j :

$$2\sigma^2 \hat{\beta} + 2\beta^2 \left[\sum_i a_i x_i - 1 \right] \sum_j x_j y_j = 0$$

or:

$$\sigma^2 \hat{\beta} + \beta^2 \left[\sum_i a_i x_i - 1 \right] \sum_j x_j y_j = 0 \quad (2)$$

Also, multiply (1) by x_j and sum over all j :

$$\begin{aligned} 2\sigma^2 \sum_j a_j x_j + 2\beta^2 \left[\sum_i a_i x_i - 1 \right] \sum_j x_j^2 &= 0 \\ \Rightarrow \sum_i a_i x_i &= \left(\frac{\beta^2 \sum_i x_i^2}{\sigma^2 + \beta^2 \sum_i x_i^2} \right) \end{aligned}$$

Substituting into (2):

$$\sigma^2 \hat{\beta} + \beta^2 \left[\frac{\beta^2 \sum_i x_i^2}{\sigma^2 + \beta^2 \sum_i x_i^2} - 1 \right] \sum_i x_i y_i = 0$$

or,

$$\sigma^2 \hat{\beta} + \beta^2 \left[\frac{\beta^2 \sum_i x_i^2 - \sigma^2 - \beta^2 \sum_i x_i^2}{\sigma^2 + \beta^2 \sum_i x_i^2} \right] \sum_i x_i y_i = 0$$

or,

$$\begin{aligned} \hat{\beta} &= \left[\frac{\beta^2 \sigma^2}{\sigma^2 + \beta^2 \sum_i x_i^2} \right] \frac{\sum_i x_i y_i}{\sigma^2} \\ &= \left(\frac{\beta^2}{\sigma^2 + \beta^2 \sum_i x_i^2} \right) \sum_i x_i y_i \\ &= \left(\frac{\beta^2 \sum_i x_i^2}{\sigma^2 + \beta^2 \sum_i x_i^2} \right) b \end{aligned}$$

where $b = (\sum_i x_i y_i / \sum_i x_i^2)$ is the OLS estimator.

As $\hat{\beta}$ depends on β and σ^2 , it can't be applied in practice.

(b) Now consider minimizing:

$$H = h \left[\frac{\text{var}(\hat{\beta})}{\sigma^2} \right] + (1-h) \left[\frac{\text{Bias}(\hat{\beta})}{\beta} \right]^2 ; 0 < h < 1$$

where again, $\hat{\beta} = \sum_i a_i y_i$. So,

$$\partial H / \partial a_j = 2ha_j + 2(1-h)x_j \left[\sum_i a_i x_i - 1 \right] = 0.$$

Multiply by y_j and sum:

$$h \sum_j a_j y_j + (1-h) \sum_j x_j y_j \left[\sum_i a_i x_i - 1 \right] = 0.$$

or

$$h\hat{\beta} + (1-h) \sum_i x_i y_i \left[\sum_i a_i x_i - 1 \right] = 0.$$

Sum over j after multiplying by x_j :

$$h \sum_j a_j x_j + (1-h) \sum_j x_j^2 \left[\sum_i a_i x_i - 1 \right] = 0$$

$$\begin{aligned} &\Rightarrow \sum_i a_i x_i \left[h + (1-h) \sum_i x_i^2 \right] = (1-h) \sum_i x_i^2 \\ &\Rightarrow \sum_i a_i x_i = \frac{(1-h) \sum_i x_i^2}{h + (1-h) \sum_i x_i^2} \end{aligned}$$

So,

$$\begin{aligned} &h\hat{\beta} + (1-h) \sum_i x_i y_i \left[\frac{(1-h) \sum_i x_i^2}{h + (1-h) \sum_i x_i^2} - 1 \right] = 0 \\ &\Rightarrow h\hat{\beta} + (1-h) \sum_i x_i^2 b \left[\frac{(1-h) \sum_i x_i^2 - h - (1-h) \sum_i x_i^2}{h + (1-h) \sum_i x_i^2} \right] = 0 \\ &\Rightarrow \hat{\beta} = \left[\frac{1}{h + (1-h) \sum_i x_i^2} \right] (1-h) \sum_i x_i^2 b \\ &\quad = \left[\frac{(1-h) \sum_i x_i^2}{h + (1-h) \sum_i x_i^2} \right] b. \end{aligned}$$

This estimator is operational in any $h \in (0, 1)$, and if $h = 0$, or $h = 1$, $\hat{\beta} = b$, or $\hat{\beta} = 0$, respectively.