

BAYESIAN ECONOMETRICS

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1. GENERAL BACKGROUND

THEMES:

- (i) Unified set of principles
- (ii) Flexible use of prior information
- (iii) Explicit decision-theoretic basis for inference.
- (iv) Estimators / tests which are "optimal"

PROBABILITY:

- (i) Classical (A priori) probability.
- (ii) Long-Run Relative Frequency.
- (iii) Subjective (personalistic) probability.

Features:

- (a) Once-and-for all events O.K.
- (b) Results conditional on information set.
- (c) Results revised with more information.
- (d) "Subjective betting-odds"

BAYES' THEOREM:

Recall: $P(B|A) = P(B \cap A) / P(A)$

$$P(A|B) = P(A \cap B) / P(B)$$

$$\Rightarrow P(B|A) = P(A|B) P(B) / P(A)$$

More generally -

$$P(B_j|A) = \frac{P(A|B_j) P(B_j)}{\sum_i P(A|B_i) P(B_i)}$$

So:

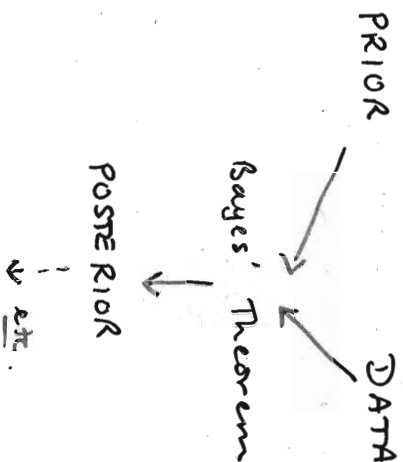
$$P(\theta|y) = \frac{P(y|\theta) P(\theta)}{P(y)}$$

$$= \frac{P(y|\theta) \cdot P(\theta)}{\int P(y|\theta) P(\theta) d\theta}$$

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So, $\int p(\theta|y) d\theta = 1$, & we can write:

$$p(\theta|y) \propto p(\theta) \cdot p(y|\theta) = p(\theta) L(\theta|y)$$



Posterior $\propto$ (Prior) $\times$ (Likelihood Fcn.)

DECISION THEORY

Loss fcn. : $L_1(\theta, \hat{\theta}) = c_1(\theta - \hat{\theta})^2$

$$L_2(\theta, \hat{\theta}) = c_2 |\theta - \hat{\theta}|$$

$$L_3(\theta, \hat{\theta}) = \begin{cases} 1 & ; |\theta - \hat{\theta}| > \epsilon \\ 0 & ; |\theta - \hat{\theta}| \leq \epsilon \end{cases}$$

(E70)

"MEL" Rule:

"Act so as to minimize posterior expected loss."

(choose $\hat{\theta}$; select a model / hypothesis)

BAYES' RULE:

"Act so as to minimize average risk."

(a) Posterior Expected Loss is:

$$\int_{-\infty}^{\infty} L(\theta, \hat{\theta}) p(\theta|y) d\theta$$

(b) Risk = Expected Loss

$$R(\hat{\theta}) = \int_y L(\theta, \hat{\theta}) p(y|\theta) dy$$

(c) Average risk

$$r(\hat{\theta}) = \int_{-\infty}^{\infty} R(\hat{\theta}) p(\theta) d\theta$$

So, Bayes' Rule $\Rightarrow$

$$\min_{(\hat{\theta})} [r(\hat{\theta})] = \min_{(\hat{\theta})} \int \int L(\theta, \hat{\theta}) p(y|\theta) dy p(\theta) d\theta$$

$$= \min_{(\hat{\theta})} \int \left[\int L(\theta, \hat{\theta}) p(\theta|y) p(y) dy \right] d\theta$$

$$= \min_{(\hat{\theta})} \int p(y) \left[\int L(\theta, \hat{\theta}) p(\theta|y) d\theta \right] dy$$

(This requires convergence of $\iint$ - Fubini)

Integrand is non-negative, so $\Leftrightarrow$

$$\min_{(\hat{\theta})} \int L(\theta, \hat{\theta}) p(\theta|y) d\theta$$

So, often, Bayes' Rule $\Rightarrow$ MEL Rule.

(If Bayes' Rule undefined, can still use MEL Rule.)

Results:

$$(a) \quad L_1 : \hat{\theta} = \text{posterior mean}$$

$$(b) \quad L_2 : \hat{\theta} = \text{posterior median}$$

$$(c) \quad L_3 : \hat{\theta} = \text{posterior mode}$$

If $p(\theta)$ is "proper", then Bayes' Rules are consistent + admissible, by construction.
(Try to prove latter result.)

Example:

$$\min_{(\hat{\theta})} \int c_1 (\theta - \hat{\theta})^2 p(\theta|y) d\theta$$

$$\Rightarrow c_1 \int \frac{d}{d\hat{\theta}} (\theta - \hat{\theta})^2 p(\theta|y) d\theta = 0$$

$$\Rightarrow \int 2(\hat{\theta} - \theta) p(\theta|y) d\theta = 0$$

$$\Rightarrow \hat{\theta} \int p(\theta|y) d\theta = \int \theta p(\theta|y) d\theta$$

$$\Rightarrow \hat{\theta} = E(\theta|y).$$

(Check 2nd. Der.)

Example:

$$X_i \sim N(\mu, \sigma_0^2) ; \sigma_0^2 \text{ known}$$

Information about unknown μ represented

$$by N(\bar{\mu}, \bar{v}).$$

$$x = (x_1, x_2, \dots, x_n)$$

$$\text{Now, } p(x|\mu) = L(\mu|x)$$

$$= \left(\frac{1}{\sigma_0 \sqrt{2\pi}}\right)^n \exp\left\{-\frac{1}{2\sigma_0^2} \sum_{i=1}^n (x_i - \mu)^2\right\}$$

$$\propto \exp\left\{-\frac{1}{2\sigma_0^2} \sum_{i=1}^n (x_i - \mu)^2\right\}$$

kernel of density.

$$p(\mu) = \frac{1}{\sqrt{2\pi\bar{v}}} \exp\left[-\frac{1}{2\bar{v}} (\mu - \bar{\mu})^2\right]$$

$$\propto \exp\left\{-\frac{1}{2\bar{v}} (\mu - \bar{\mu})^2\right\}$$

Bayes' Theorem:

$$p(\mu|x) \propto p(\mu) \cdot p(x|\mu)$$

$$\text{So, } p(\mu|x) \propto \exp\left[-\frac{1}{2} \left\{ \frac{1}{\bar{v}} (\mu - \bar{\mu})^2 + \frac{1}{\sigma_0^2} \sum_{i=1}^n (x_i - \mu)^2 \right\}\right]$$

Now,

$$\sum_{i=1}^n (x_i - \mu)^2 = \sum_{i=1}^n (x_i - \bar{x})^2 + n(\bar{x} - \mu)^2$$

So,

$$p(\mu|x) \propto \exp\left\{-\frac{1}{2} \left[\frac{1}{\bar{v}} (\mu - \bar{\mu})^2 + \frac{1}{\sigma_0^2} \sum_{i=1}^n (x_i - \bar{x})^2 \right. \right.$$

$$\left. + \frac{n}{\sigma_0^2} (\bar{x} - \mu)^2 \right]\}$$

$$\propto \exp\left\{-\frac{1}{2} \left[\mu^2 \left(\frac{1}{\bar{v}} + \frac{n}{\sigma_0^2} \right) - 2\mu \left(\frac{\bar{\mu}}{\bar{v}} + \frac{n\bar{x}}{\sigma_0^2} \right) \right. \right.$$

$$\left. + \text{constant} \right]\}$$

$$\propto \exp\left\{-\frac{1}{2} (a\mu^2 + b\mu + c)\right\}$$

$$\propto \exp\left\{-\frac{1}{2} \left[a \left(\mu + \frac{b}{2a} \right)^2 + \left(c - \frac{b^2}{4a} \right) \right] \right.$$

So:

$$p(\mu|x) \propto \exp\left\{-\frac{1}{2} \left(\frac{1}{\bar{v}} + \frac{n}{\sigma_0^2} \right) \left(\mu - \left(\frac{\frac{\bar{\mu}}{\bar{v}} + \frac{n\bar{x}}{\sigma_0^2}}{\frac{1}{\bar{v}} + \frac{n}{\sigma_0^2}} \right) \right)^2 \right.$$

$$\left. \sim N\left[\bar{\mu}, \bar{v}\right] \right\}$$

$$\text{where: } \frac{1}{\bar{v}} = \left[\frac{1}{\bar{v}} + \frac{n}{\sigma_0^2} \right]$$

$$\bar{\mu} = \frac{\left[\frac{1}{\sigma^2} \bar{\mu} + \frac{1}{\sigma^2} \bar{x} \right]}{\left[\frac{1}{\sigma^2} + \frac{1}{\sigma^2} \right]}$$

(Recall: $\bar{x} \sim N(\mu, \sigma^2/n)$)

(a) Weighted-average interpretation.

(b) $\bar{\mu}$ is BAYES estimator of μ .

(c) Example of "Natural Conjugacy".

MARGINAL DENSITIES:

Let's see how flexible the Bayesian approach is w.r.t. "nuisance parameters".

e.g., $\Theta = (\beta, \sigma) = (\beta_1, \beta_2, \dots, \beta_k, \sigma)$

$$P(\Theta|y) = P(\beta_1, \dots, \beta_k, \sigma|y)$$

$$P(\beta|y) = \int_0^\infty P(\beta, \sigma|y) d\sigma$$

$$P(\sigma|y) = \int_{-\infty}^\infty \dots \int_{-\infty}^\infty P(\beta, \sigma|y) d\beta_1 d\beta_2 \dots d\beta_k$$

Then, $\hat{\beta}_i = E(\beta_i|y)$

$$= \int_{-\infty}^\infty \beta_i \cdot P(\beta_i|y) d\beta_i$$

$$\text{where } P(\beta_i|y) = \int_{-\infty}^\infty \dots \int_{-\infty}^\infty P(\beta_1, \dots, \beta_{i-1}, \beta_i, \dots, \beta_k) d\beta_1 \dots d\beta_{i-1} d\beta_{i+1} \dots d\beta_k$$

$$\text{var.}(\beta_i|y) = \int_{-\infty}^\infty (\beta_i - \hat{\beta}_i)^2 P(\beta_i|y) d\beta_i$$

(Enables us to provide point & interval estimates of β_i , ..., etc.

Example:

$$C_t = b Y_t + u_t$$

$$u_t \sim N(0, \sigma^2) ; t=1, \dots, n.$$

Prior Information:

$$0 \leq b \leq 1$$

$$E(b) = 0.8$$

$$\text{var.}(b) \leq 0.001$$

$$0 < \sigma < \infty$$

Likelihood Function:

$$P(C|b, \sigma) \propto \sigma^{-n} \exp\left[-\frac{1}{2\sigma^2} \sum_t (C_t - bY_t)^2\right]$$

Prior pdf:

$$P(b, \sigma) = p(b) \cdot p(\sigma)$$

$$p(\sigma) \propto \sigma^{-1} \quad ; \quad 0 < \sigma < \infty$$

$$p(b) \propto b^{\alpha-1} (1-b)^{\tau-1} \quad ; \quad 0 \leq b \leq 1$$

($\alpha, \tau > 0$)

What are α and τ ?

$$E(b) = \left(\frac{\alpha}{\alpha+\tau}\right) = 0.8$$

$$\text{var}(b) = \alpha\tau / [(\alpha+\tau)^2 (\alpha+\tau+1)] \approx 0.001$$

$$\text{Solve } \Rightarrow \alpha = 127.2, \tau = 31.8$$

$$\text{So: } P(b, \sigma|C) \propto P(b, \sigma) P(C|b, \sigma)$$

$$\propto \sigma^{-(n+1)} \exp\left\{-\frac{1}{2\sigma^2} \sum_t (C_t - bY_t)^2\right\} b^{126.2} (1-b)^{30.8}$$

(a) Proportionality constant?

(b) What are $\hat{b}, \hat{\sigma}$?

(c) Practical issues?

We can obtain the marginal posterior for 'b' by integrating out 'σ' analytically (see pp. 21-22 later on):

$$P(b|C) = \int_0^\infty P(b, \sigma|C) d\sigma$$

$$\propto \int_0^\infty \sigma^{-(n+1)} \exp\left\{-\frac{1}{2\sigma^2} \sum_t (C_t - bY_t)^2\right\} \cdot b^{126.2} (1-b)^{30.8} d\sigma$$

$$\propto \left[\sum_t (C_t - bY_t)^2 \right]^{-n/2} b^{126.2} (1-b)^{30.8}$$

As $0 \leq b \leq 1$, we can easily normalize $P(b|C)$ using numerical integration, & then find the posterior mean, etc.


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* NOW OBTAIN THE MEAN OF THE (NORMALIZED) MARGINAL POSTERIOR FOR "b":
      inted b lower: upper mean=normm(b**127.2)*(1-b)**30.8)/ &
      exp(18*log(s:(1-b**2*e:2-2*b*s:3))
OBSERVATION LOWERBOUND UPPERBOUND INTEGRAL
          1           0.0000000         1.0000000        0.89310850

D mean
MEAN
      0.8931085

*
*
THE BAYES ESTIMATOR OF "b" IS VERY SIMILAR TO THE MEAN OF "b"
AS THE DATA EVIDENCE IS VERY STRONG - SEE THE t-RATIO ON THE SLOPE
COEFFICIENT IN THE OLS (MLE) REGRESSION EARLIER.
* ALSO NOTE THAT THE BAYES ESTIMATOR IS A LITTLE SMALLER THAN THE MLE,
* BECAUSE THE PRIOR MEAN WAS 0.8, WHICH WAS LESS THAN THE MLE.
NOW PLOT THE MARGINAL POSTERIOR OVER THE FULL RANGE:

set nodoccho
do #=1,101
genl x:#0-(#-1)*0.01
genl post1:=normm(x:#)**126.2)*((1-x:#)**30.8)/exp(18*log(s:l+x:#)**2*&
s:-2-2*x:#*e:3))
endo
***** EXECUTION BEGINNING FOR DO LOOP # = 101
**** EXECUTION FINISHED FOR DO LOOP # = 101
**** 1 101

plot post1 x/nopretty

REQUIRED MEMORY IS PAR= 104 CURRENT PAR= 300000
101 OBSERVATIONS

**POST1
M-MULTIPLE POINT
*

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	0.000	0.250	0.500	0.750	1.000
X					

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* THIS PLOT IS NOT VERY INFORMATIVE, SO REFINIE THE SCALE AND "ZOOM IN"
* ON THE RELEVANT RANGE OF VALUES FOR "p":
*
* set nodeachs
* do #=1,201
*   genl post2:#=nomr*((1xx:#)+*126.2)*((1-1xx:#)**30.8)/exp(18*Log(1+1*(1xx:#)**2)&
*   s:2-2*(1xx:#)**3))
*   ***** EXECUTION BEGINNING FOR DO LOOP # = 1
*   ***** EXECUTION FINISHED FOR DO LOOP # = 201
*   _cmpl 1 201
*
* _plot post2 xx
*
REQUIRED MEMORY IS PAR= 105 CURRENT PAR= 300000
201 OBSERVATIONS
*-----*
*POST2
M=MULTIPLE POINT
86.0000
75.7659
71.5779
67.3688
63.1588
58.9497
54.7397
50.5266
46.3166
42.1055
37.8955
33.6844
29.4744
25.2644
21.0543
16.8442
12.6322
8.4211
4.2105
0.0000

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PROS & CONS OF BAYESIAN ANALYSIS:

- (a) Unity of approach to all types of inference.
Prior; L.F.; loss fn.; Bayes' Theorem;
MEL Rule.
- (b) Prior information incorporated explicitly
and flexibly.
- (c) Forces declaration of prior information.
- (d) Explicit use of decision theory.
- (e) Admissible estimators & tests.
- (f) Nuisance parameters handled easily.

But:

- (a) Difficulty / cost of specifying prior.
- (b) Integration / computational issues.