

ECON 546  
Answers to Supplementary Problems, III

Q.1.  $L = \theta^{\sum x_i} (1-\theta)^{\sum (1-x_i)}$

$$\log L = \sum_i x_i \log \theta + \sum_i (1-x_i) \log (1-\theta)$$

$$\frac{\partial \log L}{\partial \theta} = n\bar{x}/\theta + \frac{n(1-\bar{x})}{(1-\theta)} (-1) = 0$$

$$\Rightarrow n\bar{x}(1-\hat{\theta}) = n(1-\bar{x})\hat{\theta}$$

$$\Rightarrow \bar{x} = \hat{\theta}$$

$$\Rightarrow \hat{\theta} = \bar{x}.$$

$$(\frac{\partial^2 \log L}{\partial \theta^2}) = -\frac{n\bar{x}}{\theta^2} - (-1)(-1) \frac{n(1-\bar{x})}{(1-\theta)^2}$$

$$= -\frac{n\bar{x}}{\theta^2} - \frac{n(1-\bar{x})}{(1-\theta)^2}$$

& when  $\theta = \bar{x}$ :  $(\frac{\partial^2 \log L}{\partial \theta^2}) = -\frac{n\bar{x}}{\bar{x}^2} - \frac{n(1-\bar{x})}{(1-\bar{x})^2}$

$$= -\frac{n}{\bar{x}} - \frac{n}{(1-\bar{x})},$$

more generally, note that for any  $\theta$ :

$$(\frac{\partial^2 \log L}{\partial \theta^2}) = \frac{-n(1-\theta)^2 - n(1-\bar{x})\theta^2}{(1-\theta)^2 \theta^2}$$

& the  $x_i$ 's must lie in  $[0, 0.5]$ , so  $(1-\bar{x}) > 0$ .

This implies  $(\frac{\partial^2 \log L}{\partial \theta^2}) < 0$ , everywhere.

(2)

$$\begin{aligned}
 \text{(a)} \quad \tilde{\log L}_u &= \sum_i x_i \log \hat{\theta} + \sum_i (1-x_i) \log (1-\hat{\theta}) \\
 &= n\bar{x} \log \bar{x} + n(1-\bar{x}) \log (1-\bar{x})
 \end{aligned}$$

$$\begin{aligned}
 \tilde{\log L}_R &= n\bar{x} \log_e(0.25) + n(1-\bar{x}) \log_e(0.75) \\
 &= n\bar{x} (-1.3863) + (-0.288)n(1-\bar{x}) \\
 &= (-1.0986n\bar{x} - 0.288n)
 \end{aligned}$$

$$LRT = -2 \log (\tilde{L}_R / \tilde{L}_u) = 2 [\tilde{\log L}_u - \tilde{\log L}_R]$$

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$$I(\theta) = -E[\partial^2 \log L / \partial \theta^2] = E\left[\frac{n\bar{x}}{\theta^2} + \frac{n(1-\bar{x})}{(1-\theta)^2}\right]$$

$$\& E(x_i) = \theta. \quad (\Rightarrow E(\bar{x}) = \theta)$$

$$\begin{aligned}
 \text{So, } I(\theta) &= (n\theta/\theta^2) + n(1-\theta)/(1-\theta)^2 \\
 &= (n/\theta) + n/(1-\theta)
 \end{aligned}$$

$$IA(\theta) = \lim_{n \rightarrow \infty} [{}^n I(\theta)] = (1/\theta - \frac{1}{1-\theta})$$

$$\begin{aligned}
 I^* \text{ is such that } \lim [{}^n I^*] &= IA, \text{ & so} \\
 I^* &= (1/\hat{\theta} + 1/(1-\hat{\theta})).
 \end{aligned}$$

$$\begin{aligned}
 W &= (\hat{\theta} - 1/4)^2 / \left( \frac{n}{\hat{\theta}} + \frac{n}{1-\hat{\theta}} \right)^{-1} = \left[ \frac{(\hat{\theta} - 1/4)^2}{\left( \frac{n(1-\hat{\theta}) + n\hat{\theta}}{\hat{\theta}(1-\hat{\theta})} \right)^{-1}} \right] \\
 &= \frac{(\hat{\theta} - 1/4)^2}{\hat{\theta}(1-\hat{\theta})} (n) * = \frac{n(1-\hat{\theta})}{\hat{\theta}} \frac{n(1-\bar{x})}{\bar{x}}
 \end{aligned}$$

(3)

$$\begin{aligned}
 LM &= \left[ \frac{n\bar{x}}{\theta} - \frac{n(1-\bar{x})}{1-\theta} \right]^2 \Big|_{\theta=0.25} \Big|_{\theta=0.25} \bar{I}^{\bar{x}-1} \\
 &= [4n\bar{x} - 4n(1-\bar{x})/3]^2 (4n + 4n/3)^{-1} \\
 &= 16n^2 (\bar{x} - (1-\bar{x})/3)^2 / \{ 4n(4/3) \} \\
 &= 3n (\bar{x} - (1-\bar{x})/3)^2.
 \end{aligned}$$

(b) Reject  $H_0$  if test statistic  $> 2.71$ . Assuming null distribution is  $\chi^2_{(1)}$ , which will be valid asymptotically.

Q. 2 (a)  $L = \theta^n \prod_i y_i^{\theta-1}; \log L = n \log \theta + (\theta-1) \sum_i \log y_i$

$$\begin{aligned}
 (\partial \log L / \partial \theta) &= n/\theta + \sum_i \log y_i = 0 \\
 \Rightarrow \tilde{\theta} &= -n / \sum_i \log y_i
 \end{aligned}$$

[ Note that  $0 < y_i < 1 \Rightarrow \log y_i < 0 \Rightarrow \tilde{\theta} > 0$  ]

$$(\partial^2 \log L / \partial \theta^2) = -n/\theta^2; < 0 \text{ everywhere}$$

$$\log \tilde{L}_u = n \log \tilde{\theta} + (\tilde{\theta}-1) \sum_i \log y_i$$

$$\log \tilde{L}_R = 0 + 0 = 0$$

$$\begin{aligned}
 \text{So, } LRT &= 2 [\log \tilde{L}_u - \log \tilde{L}_R] \\
 &= 2n \log \tilde{\theta} + 2(\tilde{\theta}-1) \sum_i \log y_i \quad *
 \end{aligned}$$

(4)

$$I(\theta) = -E[\partial^2 \log L / \partial \theta^2] = 1/\theta^2$$

$$IA(\theta) = \lim_{n \rightarrow \infty} [ \frac{1}{n} I(\theta) ] = 1/\theta^2$$

$$I^* = 1/\hat{\theta}^2$$

$$\omega = (\hat{\theta}-1)^2 (1/\hat{\theta}^2).$$

$$\text{LM} = \frac{(n + \sum_i \log y_i)^2}{n} \quad (\cancel{\hat{\theta}^2})$$

(b) As for Q. 1. (a), but critical value = 3.84.

$$Q=3. L = (\theta/2\pi)^{n/2} e^{n\theta} (\prod_i y_i^{-3/2}) \exp \left[ -\frac{\theta}{2} \sum_i (y_i + y_i^{-1}) \right]$$

$$\log L = \frac{n}{2} \log \theta - \frac{n}{2} \log 2\pi + n\theta + \sum_i \log (y_i^{-3/2})$$

$$- (\theta/2) \sum_i (y_i + y_i^{-1})$$

$$= \frac{n}{2} \log \theta - \frac{n}{2} \log 2\pi + n\theta - \frac{3}{2} \sum_i \log y_i$$

$$- (\theta/2) \sum_i (y_i + y_i^{-1})$$

$$\frac{\partial \log L}{\partial \theta} = \frac{n}{2\theta} + n - \frac{1}{2} \sum_i (y_i + y_i^{-1}) = 0$$

$$\Rightarrow \frac{n}{2\hat{\theta}} = \left[ \frac{1}{2} \sum_i (y_i + y_i^{-1}) - n \right]$$

$$\Rightarrow \hat{\theta} = \left( \frac{n}{2} \left[ \frac{1}{2} \sum_i (y_i + y_i^{-1}) - n \right] \right)^{-1}$$

(5)

$$(\partial^2 \log L / \partial \theta^2) = -\frac{1}{2\theta^2} < 0$$

$$I = n/2\theta^2$$

$$IA = \frac{1}{2}\theta^2 ; I^* = \frac{1}{2}\tilde{\theta}^2$$

$$\log \tilde{L}_R = -\frac{1}{2} \log 2\pi + n - \frac{3}{2} \sum_i \log y_i - \frac{1}{2} \sum_i (y_i + y_i^{-1})$$

$$\log \tilde{L}_u = \frac{1}{2} \log \tilde{\theta} - \frac{1}{2} \log 2\pi + n \tilde{\theta} - \frac{3}{2} \sum_i \log y_i - (\tilde{\theta}/2) \sum_i (y_i + y_i^{-1})$$

$$\text{So, LRT} = 2 \left[ \frac{1}{2} \log \tilde{\theta} + n(\tilde{\theta} - 1) - \sum_i (y_i + y_i^{-1}) [\tilde{\theta} - 1]/2 \right]$$

$$\omega = n(\tilde{\theta} - 1)^2 / 2\tilde{\theta}^2$$

$$\begin{aligned} LM &= \left[ \frac{1}{2} + n - \frac{1}{2} \sum_i (y_i + y_i^{-1}) \right]^2 \left( \frac{2}{n} \right) \\ &= [1 + 2 - \frac{1}{n} \sum_i (y_i + y_i^{-1})]^2 \\ &= [3 - \frac{1}{n} \sum_i (y_i + y_i^{-1})]^2. \end{aligned}$$

$$\underline{\text{Q.4.}} \quad \lambda = \left( \frac{1}{\prod y_i!} \right) e^{-n\lambda} \lambda^{\sum y_i}$$

$$\log L = - \sum_i \log(y_i!) - n\lambda + n\bar{y} \log \lambda$$

$$(\partial \log L / \partial \lambda) = -n + \bar{y}n/\lambda = 0$$

$$\Rightarrow \bar{\lambda} = \bar{y}.$$

(6)

$$(\partial^2 \log L / \partial \lambda^2) = -n\bar{y}/\lambda^2 < 0.$$

$$I = n\lambda^2 E(\bar{y}) = (n\lambda/\lambda^2) = n/\lambda.$$

$$(E(y_i) = \lambda))$$

$$\text{So, } IA = 1/\lambda \quad \& \quad I^* = n/\bar{y} = (n/\bar{y}).$$

$$\underline{\text{a)}} \quad LRT = 2 [\text{const} - n\bar{y} + n\bar{y} \log \bar{y} - \text{const} + 5n - n\bar{y} \log 5]$$

$$= 2 [n\bar{y} (\log \bar{y} - 1 - \log_2 5) + 5n]$$

$$= 2 [n\bar{y} (\log \bar{y} - 2 \cdot 609) + 5n],$$

$$W = n(\bar{y} - 5)^2/\bar{y}$$

$$LM = 5[(n\bar{y}/5) - n]^2/n$$

$$\underline{\text{b)}} \quad LRT = 2 [2300 (\log(4.6) - 2 \cdot 609) + 2500] \quad (\bar{y} = 4.6)$$

$$= 18.476.$$

$$W = 17.391$$

$$LM = 16$$

(Easily reject  $H_0$  at any reasonable sig. level.)

(7)

$$\underline{Q.5.} \quad p(x, y) = \theta \exp[-(\beta + \theta)y] [(\beta y)^x] / x!$$

$$(a) L = \theta^n \exp[-(\theta + \beta) \sum y_i] \prod (\beta y_i^{x_i}) / \prod y_i!$$

$$\propto \theta^n \exp[-(\beta + \theta) n \bar{y}] \beta^{\sum x_i}$$

$$\log L = c + n \log \theta - (\beta + \theta) n \bar{y} + n \bar{x} \log \beta$$

$$(\partial \log L / \partial \theta) = n/\theta - n\bar{y} = 0 \Rightarrow \hat{\theta} = \bar{y}.$$

$$(\partial \log L / \partial \beta) = (n\bar{x}/\beta) - n\bar{y} \Rightarrow \hat{\beta} = (\bar{x}/\bar{y})$$

$$(\partial^2 \log L / \partial \theta^2) = -n/\theta^2; \quad (\partial^2 \log L / \partial \beta^2) = -n\bar{x}/\beta^2$$

$$(\partial^2 \log L / \partial \theta \partial \beta) = 0$$

$$H = \begin{bmatrix} -n/\hat{\theta}^2 & 0 \\ 0 & -n\bar{x}/\hat{\beta}^2 \end{bmatrix}$$

(negative-definite as co-factors have signs, -, +)

So  $(\hat{\theta}, \hat{\beta})$  locates a max. of  $\log L$ .

(b) By invariance MLE of  $\theta/(\theta + \beta)$  is

$$[\hat{\theta}/(\hat{\theta} + \hat{\beta})] = [\bar{y}/(\bar{y} + \bar{x}/\bar{y})] = \bar{y}^2/(\bar{x} + \bar{y}^2).$$

$$(c) \quad p(x) = \int_0^\infty p(x, y) dy = \int_0^\infty \theta e^{-(\beta + \theta)y} \beta^x y^x / x! dy$$

$$= \left( \frac{\theta \beta^x}{x!} \right) \int_0^\infty e^{-(\beta + \theta)y} y^x dy$$

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$$\text{Let } t = (\beta + \alpha)y ; dy = (\beta + \alpha)^{-1} dt$$

$$P(x) = \left( \frac{\alpha \beta^x}{x!} \right) \int_0^\infty e^{-t} \left( \frac{t}{\beta + \alpha} \right)^x \left( \frac{1}{\beta + \alpha} \right) dt$$

$$= \frac{\alpha \beta^x}{x! (\beta + \alpha)^{x+1}} \int_0^\infty e^{-t} t^x dt$$

$$= \frac{\alpha \beta^x}{x! (\beta + \alpha)^{x+1}} \int_0^\infty e^{-t} t^{(x+1)-1} dt$$

$$= \frac{\alpha \beta^x \Gamma(x+1)}{x! (\beta + \alpha)^{x+1}} = \frac{\alpha \beta^x x!}{x! (\beta + \alpha)^{x+1}}$$

$$= \alpha \beta^x / (\beta + \alpha)^{x+1} ; x = 0, 1, 2, \dots$$

$$= \delta (1-\delta)^x$$

$$\text{where } \delta = \frac{1}{\beta + \alpha} ; \gamma = \left(1 - \frac{\alpha \beta}{\beta + \alpha}\right)$$

$$\text{By invariance, } \tilde{\delta} = \left(1 - \frac{\widehat{\alpha} \widehat{\beta}}{\widehat{\beta} + \widehat{\alpha}}\right)$$

$$\text{or, } \tilde{\delta} = \left( \frac{\widehat{\beta} + \widehat{\alpha} - \widehat{\alpha} \widehat{\beta}}{\widehat{\beta} + \widehat{\alpha}} \right) = \left( \frac{\bar{x}/\bar{y} + \bar{y} - \bar{x}}{\bar{x}/\bar{y} + \bar{y}} \right)$$

$$= (\bar{x} + \bar{y}^2 - \bar{x}\bar{y}) / (\bar{x} + \bar{y}^2)$$

(9)

(d) As  $H$  is diagonal, we can just focus on the  $(2, 2)$  element:  $-n\bar{x}/\beta^2$ .

$$\begin{aligned} LRT &= 2 \left[ c + n \log \bar{y} - (\bar{y} + \bar{x}/\bar{y}) n \bar{y} + n \bar{x} \log(\bar{x}\bar{y}) \right. \\ &\quad \left. - c - n \log(1) + \bar{y} n \bar{y} - n \bar{x} \log(1) \right] \\ &= 2 \left[ n \log \bar{y} - n \bar{y}^2 + n \bar{y} \bar{x} + n \bar{x} \log \bar{x} - n \bar{x} \log \bar{y} \right. \\ &\quad \left. + \bar{y}^2 n \right] \\ &= 2 \left[ n \bar{x} \bar{y} + n \log \bar{y} (1 - \bar{x}) + n \bar{x} \log \bar{x} \right] \end{aligned}$$

Reject  $H_0$  if  $LRT > \chi_{c,1}^2$ , critical value.