

Supplementary Exercises V - Solution

$$\text{Q-1. } L \propto e^{-n\lambda} \lambda^{\sum x_i} = e^{-n\lambda} \lambda^{n\bar{x}}$$

(a)

$$p(\lambda) \propto \lambda^{\gamma-1} e^{-\lambda}$$

$$\text{So, } p(\lambda|x) \propto \lambda^{n\bar{x}+\gamma-1} e^{-(n+1)\lambda}$$

& this is Gamma, with parameters $(n\bar{x}+\gamma)$ and $(n+1)$. The prior had parameters γ and $1'$.

(b) The mean of a Gamma density is γ (or in the case of the posterior, $(n\bar{x}+\gamma)$). And this is the Bayes estimator under quadratic loss. That is $\hat{\lambda} = (n\bar{x}+\gamma)$.

Mode:

$\frac{\partial p(\lambda|x)}{\partial \lambda}$ will be zero when $(\frac{\partial \log p}{\partial \lambda})=0$.

$$\begin{aligned} \log[p(\lambda|x)] &= \text{const.} + (n\bar{x}+\gamma-1) \log \lambda \\ &\quad - (n+1)\lambda. \end{aligned}$$

$$(\frac{\partial \log p}{\partial \lambda}) = (n\bar{x}+\gamma-1)/\lambda - (n+1) = 0$$

$$\Rightarrow \lambda^* = \left(\frac{n\bar{x}+\gamma-1}{n+1} \right)$$

This is the Bayes' estimator under 0-1 loss.

Now the data are Poisson & mean = λ .

$$\text{So, } E(\hat{\lambda}) = nE(\bar{x})+\gamma = n\lambda+\gamma \neq \lambda. \quad (\text{Biased})$$

$$\text{But } E(\lambda^*) = \left(\frac{nE(\bar{x})+\gamma-1}{n+1} \right) = \left(\frac{n\lambda+\gamma-1}{n+1} \right) \neq \lambda.$$

(2)

Q.2.

$$\text{(a)} \quad L(\theta|y) = \theta^n e^{-\theta \sum y_i}$$

$$p(\theta) = \frac{\theta^{\alpha-1} e^{-\theta/\beta}}{\beta^\alpha \Gamma(\alpha)}$$

By Bayes' Theorem,

$$p(\theta|y) \propto \frac{\theta^{n+\alpha-1} e^{-\theta[\frac{1}{\beta} + \sum y_i]}}{\beta^\alpha \Gamma(\alpha)}$$

$$\propto \theta^{(n+\alpha)-1} e^{-\theta[\frac{1}{\beta} + \sum y_i]^{-1}}$$

which is a Gamma density, with parameters $(n+\alpha)$ and $[\frac{1}{\beta} + \sum y_i]^{-1}$, as required.

(b) $E(\theta|y) = (n+\alpha)/[\frac{1}{\beta} + \sum y_i]$, & this is the Bayes estimator of θ under quadratic loss.

(c) The mode is at $(n+\alpha-1)/[\frac{1}{\beta} + \sum y_i]$, & this is Bayes estimator of θ if $n+\alpha > 1$.

Q.3.

$$\text{(a)} \quad \hat{\beta} = (A + X'X)^{-1}(A\bar{\beta} + X'y)$$

$$\text{where } \begin{cases} p(\beta|\sigma) = N[\bar{\beta}, \sigma^2 A^{-1}] \\ p(\sigma) = \text{Gamma}(\omega, c). \end{cases}$$

The model is $y = X\beta + \varepsilon$; $\varepsilon \sim N[0, \sigma^2 I]$.

(3)

$$\text{So: } E(\hat{\beta}) = (A + X'X)^{-1} (A\bar{\beta} + X'E(y))$$

$$= (A + X'X)^{-1} (A\bar{\beta} + X'X\beta)$$

$$\neq \beta.$$

So, the estimator is biased.

(b) Bayesians aren't worried about performance in "repeated samples", or "on average".

(c) The MLE for β is OLS = $b = (X'X)^{-1}X'y$,
 $\& V(b) = \sigma^2(X'X)^{-1} \propto (X'X)^{-1}$. So, here we
 are going to choose $\sigma^2 A^{-1} = \sigma^2 (X'X)^{-1}$.

$$\begin{aligned}\text{Then, } \hat{\beta} &= ([\sigma^2(X'X)^{-1}]^{-1} + (X'X))^{-1} \\ &\quad \cdot ([\sigma^2(X'X)^{-1}]^{-1} \bar{\beta} + X'y) \\ &= [\frac{1}{\sigma^2} X'X + X'X]^{-1} [\frac{1}{\sigma^2}(X'X)\bar{\beta} + X'y] \\ &= [\frac{1}{\sigma^2} X'X + X'X]^{-1} [\cancel{\frac{1}{\sigma^2}}(X'X)\bar{\beta} + (X'X)b] \\ &= [\frac{1}{\sigma^2} + 1]^{-1} (X'X)^{-1} (X'X) [\frac{1}{\sigma^2} \bar{\beta} + b] \\ &= (\frac{1}{\sigma^2} \bar{\beta} + b) / (\frac{1}{\sigma^2} + 1)\end{aligned}$$

$$\& E(\hat{\beta}) = (\frac{1}{\sigma^2} \bar{\beta} + \beta) / (\frac{1}{\sigma^2} + 1). \quad \#.$$

(d) It involves seeing the sample of X data before constructing the prior.

(4)

$$\text{Q.4 (a)} \quad p(x | \mu_0, \sigma) = \frac{1}{\sigma \sqrt{2\pi}} \exp\left[-\frac{1}{2\sigma^2}(x - \mu_0)^2\right]$$

$$\propto \sigma^{-1} \exp\left[-\frac{1}{2}\sigma^{-2}(x - \mu_0)^2\right].$$

Let $\tau = \sigma^{-2}$:

$$p(x | \mu_0, \tau) \propto \tau^{1/2} \exp\left[-\frac{1}{2}\tau(x - \mu_0)^2\right].$$

Let $p(\tau) = \text{Gamma}(\alpha, \sigma)$

$$\propto e^{-\alpha\tau} \tau^{\alpha-1}$$

Apply Bayes' Theorem:

$$p(\tau | x) \propto e^{-\alpha\tau} \tau^{\alpha-1} \tau^{n/2} \exp\left[-\frac{\tau}{2} \sum_i (x_i - \mu_0)^2\right]$$

$$\propto \tau^{(\alpha+n/2)-1} \exp[-\delta\tau + \alpha\tau]$$

$$\text{where } \delta = \frac{1}{2} \sum_i (x_i - \mu_0)^2 \quad (\text{known.})$$

$$\text{So, } p(\tau | x) \propto \tau^{(\alpha+n/2)-1} e^{-\delta\tau}$$

$$\text{where } \delta = \alpha + \alpha.$$

This is Gamma $[(\alpha+n/2), \alpha + \frac{1}{2} \sum_i (x_i - \mu_0)^2]$

so we have Conjugacy.

(5)

(b) $E(\bar{\tau}|x) = (\delta + \frac{1}{n})$, & this is the Bayes' estimator under quadratic loss.

Following from the answer to Q. 2(b), the mode of $P(\bar{\tau}|x)$ is at:

$$\bar{\tau}^* = \left[\frac{\delta + \frac{1}{n} - 1}{\delta + \frac{1}{n} \sum (x_i - \mu_0)^2} \right]$$

(c) Bayes' estimators are consistent. So, in the first case, a consistent estimator of τ is $(\delta + \frac{1}{n})$ & for σ^2 it is $(\delta + \frac{1}{n})^{-2}$. In the second case, we'd have a consistent estimator of σ^2 by using

$$[\delta + \frac{1}{n} \sum (x_i - \mu_0)^2]^{-1} / (\delta + \frac{1}{n} - 1)^{-2}.$$