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Role of Reading Proficiency in Assessing Mathematics and Science Learning for Students from English and Non-English Backgrounds: An International Perspective

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Comparability of Mathematics and Science Scores for Students from English and Non-English Backgrounds in Australia, Canada, the UK, and the US


#### Abstract

The purpose of this research is to examine the comparability of mathematics and science


 scores for students from English language backgrounds (ELB) and those from non-English language backgrounds (NELB). We examine the relationship between English reading proficiency and performance on mathematics and science assessments and how this relationship affects comparability of scores for ELB and NELB. The research uses international assessment data and examines this relationship in four countries with English language education systems: Australia, Canada, the United Kingdom, and the United States. The findings indicate a strong relationship between reading proficiency and performance on mathematics and science assessments with reading proficiency accounting for large proportions of variance in both mathematics (up to $43 \%$ ) and science (up to $79 \%$ ) scores. In all comparisons, ELB students either outperformed NELB students or performed at the same level. However, when statistical adjustments were made for reading proficiency, in mathematics, the score gap between the groups remained in the US only, whereas the differences between the two groups became significant with higher scores for NELB in Canada. In science, the differences between NELB and ELB remained significant only in Australia. These findings point to differences in score meaning and limitations in comparing performance on mathematics and science assessments for NELB and ELB.Keywords: reading proficiency, mathematics assessment, science assessment, language backgrounds, language effects, international comparisons, ELL

Education systems around the world are faced with educating children who come from multiple language and cultural backgrounds. Typically, children from a different language and cultural background than the host country tend to have lower achievement levels on large-scale assessments. This results in an equity and fairness problem that needs to be addressed (Au, 2013; Ercikan, et al., in press; Nguyen \& Cortes, 2013; Vale et al., 2013). Differences in performance on assessments can be due to differences in achievement levels or inaccuracies in measurement of knowledge and competencies and limitations in interpretation of scores from such measurement. In mathematics and science assessments, scores are expected to indicate students' knowledge and skills in these areas. Validity of such score interpretations depends on the degree to which performance on assessments are accurate indicators of students' competencies (Kane, 2013). There are two key sources of potential threats to validity of score interpretations: construct-underrepresentation and construct-irrelevant variance (Messick, 1989).

Construct-underrepresentation can occur when a test does not provide a full representation of the targeted construct, jeopardizing the generalizability of the score inferences to the larger domain. This may occur when students have limited language proficiency in the test language by limiting their access to their knowledge and ability to respond to the items. As a result, scores are underestimated and fail to represent students' proficiency in the domain. Construct-irrelevant variance occurs when tests require competencies that are not targeted by the test, such as linguistic demands of items, cultural references, and context and format of items that may not be familiar to students. Construct-irrelevant variance also results in the underestimation of scores for students disadvantaged by linguistic and cultural requirements. In this paper we
focus on two questions that arise when these sources of threats to validity occur. To what extent are mathematics and science scores underestimated when students have limited proficiency in the language of the test? Furthermore, to what extent can scores be compared for students who have different proficiency levels in the language of the test?

## Language Background and Performance on Assessments

There is growing evidence that limited English proficiency has significant implications for students' success in mathematics and science assessments (Abedi, 2004; Abedi, Hofstetter, \& Lord, 2004; Abedi \& Lord, 2001; Butler, Bailey, Stevens, Huang, \& Lord, 2004; Kopriva, Gabel, \& Cameron, 2011; Luykx et al., 2007; Noble et al., in press; Penfield \& Lee, 2010; Solano-Flores \& Trumbull, 2003). Such research has demonstrated that English language learners (ELLs) in the US confront substantial challenges with science and mathematics assessments because of linguistically and culturally dependent content and representations in assessments. Researchers investigating the validity of score interpretations for ELL and non-ELL students demonstrated that linguistic complexity of items was associated with the identification of differential item functioning (Martiniello, 2008). In this research, linguistic features that create comprehension difficulties relate to complex vocabulary and sentence structure, including multiple clauses, long noun phrases and vocabulary. Inaccuracies in measurement are expected for all examinees. However, Noble et al. (in press) have shown that ELL students with the required knowledge and skills were more likely to respond incorrectly to a set of science assessment tasks compared to non-ELL students leading to greater measurement inaccuracy for ELL students.

To minimize the effects of limited language proficiency on assessments for ELL in the United States, accommodations including the provision of dictionaries or pop-up bilingual glossaries are often provided to reduce language complexity (See Abedi, in press and Lane, in press for comprehensive reviews of accommodations and modifications). Research on modifications has demonstrated that reducing the linguistic complexity of mathematics items resulted in higher performance for ELL students (Abedi, Hofstetter, \& Lord, 2004; Abedi \& Lord, 2001). Other research demonstrates that the gap in achievement between ELL and non-ELL students is largest when language demand is high and it is smaller for science and mathematics problem solving than for reading and writing (Abedi, Leon, \& Mirocha, 2003). When students are asked to perform mathematics computation in which linguistic demands are negligible the gap almost disappeared.

Although it is clear that the language competencies of examinees can affect their performances on assessments, this process may be far more complex than just the language competency of the examinee (Solano-Flores, 2008). Student performance on assessments is a result of a complex interaction among factors such as the student's home language and culture, the context of the test, and the language proficiencies of students. Thus, ELL students may face challenges with some types of mathematics problems even when language demand is low because notations often vary between cultures, languages, and countries. For instance, a point or a comma can represent a decimal number, depending on the cultural context. Another example is the way in which some ordinal numbers are represented. For instance, fifth is numerically represented by " 5 th" in the US, whereas Latin American countries use the " $o$ " superscript in
place of the "th" and French uses "ième." This problem may be further complicated because the symbol for degree $\left({ }^{\circ}\right)$ appears to be very similar to the Latin American notation. Moreover, in a study examining how elementary school students' prior linguistic and cultural knowledge mediates responses to science assessments, science terms were frequently interpreted in reference to everyday meanings rather than specialized scientific meanings (Luykx et al., 2007).

Competency in a language involves four language modalities-listening and reading (both receptive) and speaking and writing (productive use of language). In computer administered large-scale assessment contexts, all four of these language modalities may be utilized and may affect student performance. In paper-and-pencil large-scale assessment contexts that rely on the examinee's ability to read and understand the test questions and then respond in writing, the most relevant aspects of language competency are reading and writing proficiencies... In this research, we focus on reading competency instead of other language modalities for two reasons. The first reason is difficulty in obtaining data on language proficiency of students in all four modalities. Second, we assume competency in reading to be the most relevant language modality in paper-and-pencil multiple-choice tests, which continue to be the dominant assessment mode in large-scale assessments.

## Purpose

Previous research provides strong evidence of language effects on ELL student performance on mathematics and science assessments and threats to validity of score meaning for these students (Abedi, 2004; Abedi et al., 2004; Abedi \& Lord, 2001; Butler et al., 2004; Kopriva et al., 2011; Luykx et al., 2007; Noble et al., in press; Penfield \& Lee, 2010;

Solano-Flores \& Trumbull, 2003). The purpose of this research is to estimate a measure of this effect on mathematics and science scores and to examine the consistency of score meaning for students from English and non-English backgrounds. This effect is investigated in four countries with English language education systems: Australia, Canada, the United Kingdom (UK), and the United States (US). The Programme for International Student assessment (PISA) 2009 is the data source for the research.

## Method

The sections below describe PISA 2009 measures, samples, language groups, differential item functioning (DIF) analyses, score scale creation and analyses of covariance (ANCOVA) conducted in the research.

Measures

In 2009, PISA was administered in 65 countries/jurisdictions in 42 languages (OECD, 2010a). PISA is administered to 15 -year-old students in three-year cycles, with each cycle focusing primarily on one cognitive domain. In 2009, the major cognitive domain was reading literacy with science and mathematics as the minor domains. Our research used data on student performance on all three of these cognitive domains. In addition to these assessments, PISA collects data through background questionnaires on contexts of education from students, their parents, and school administrators. The background questionnaires include questions about home context, parental background, and student interests in and attitudes toward reading. PISA is designed to assess students' abilities to use their knowledge and skills to confront real-life challenges rather than to assess the extent to which they master specific school curriculum
(OECD, 2010a). The three domains of reading literacy, scientific literacy, and mathematics literacy are defined by PISA as follows:

Reading literacy: An individual's capacity to understand, use, reflect on and engage with written texts, in order to achieve one's goals, to develop one's knowledge and potential and to participate in society;

Scientific literacy: An individual's scientific knowledge and use of that knowledge to identify questions, to acquire new knowledge, to explain scientific phenomena, and to draw evidence based conclusions about science-related issues, understanding of the characteristic features of science as a form of human knowledge and enquiry, awareness of how science and technology shape our material, intellectual, and cultural environments, and willingness to engage in science-related issues, and with the ideas of science, as a reflective citizen;

Mathematical literacy: An individual's capacity to identify and understand the role that mathematics plays in the world, to make well-founded judgments and to use and engage with mathematics in ways that meet the needs of that individual's life as a constructive, concerned and reflective citizen. (OECD, 2010a, p. 23)

The assessment was administered to students in 13 test booklets. Each booklet contained four out of 13 possible clusters including seven reading clusters, and three clusters from each of science and mathematics. Students were randomly assigned to one of the 13 booklets to be completed within a two hour period.

The Reading Literacy Measure

For PISA 2009, countries were offered the option of administering a digital reading assessment in addition to paper-and-pencil format (Mendelovits, Ramalingam, \& Lumley, 2012). This research used response data from the paper-and-pencil format only because all four countries included in this study administered the paper and pencil format whereas only one (Australia) also administered the digital reading assessment. The reading literacy items were arranged in units that focused on a topic and included passages of texts, tables, graphs, and diagrams. There were 37 reading units containing a total of 131 reading items ( 37 items used in prior PISA assessments and 94 new items). Item formats were multiple-choice (MC, 52 items), complex multiple choice (CMC, 10 items), closed constructed response items (CCR, 13 items), short closed-constructed (CC, 11 items) and open-constructed responses (OCR, 45 items). For the CMC items, respondents were asked a series of questions from which they chose a series of answers (OECD, 2010a). Scale scores were reported for overall reading, the three reading aspects (access and retrieve, integrate and interpret, and reflect and evaluate), and two of the reading text formats (continuous text and non-continuous text).

## The Mathematics Literacy Measure

PISA 2009 mathematics literacy emphasized students' ability to formulate, solve and interpret mathematical problems in real-life situations. The assessment framework consists of three primary components that include the contexts or situations for the use of mathematics (personal, public, occupational, educational, scientific, intra-mathematical), mathematical concepts or content areas (space and shape, change and relationships, quantity, and uncertainty) and cognitive mathematical competencies used to solve problems. There were 35 mathematics
items ( $9 \mathrm{MC}, 7 \mathrm{CMC}, 3 \mathrm{CCR}, 8 \mathrm{CC}, 8 \mathrm{CCR}$ ) contained in 24 units in the PISA 2009 assessment. The mathematics assessment results were reported as a single overall mathematics scale (OECD, 2010a, 2010b).

## The Scientific Literacy Measure

The PISA 2009 scientific literacy assessment framework centered on students' science competence, knowledge and attitudes situated within contexts relevant to their everyday lives. The test items required students to apply science knowledge and use science competencies in particular contexts such as personal, social, or global contexts. Scientific competencies included identifying scientific issues, explaining phenomena scientifically, and drawing conclusions based on evidence. Scientific knowledge included both knowledge of the natural world (physics, chemistry, biological science, earth and space science and science-based technology) and knowledge about science (i.e., processes of scientific enquiry and scientific explanation). Similar to the reading items, the PISA 2009 science items were arranged in units that provided a common stimulus and established the context for the items. A variety of stimuli were used such as passages of text, photographs, tables, graphs, and diagrams. Most units assessed more than one scientific competency and more than one knowledge category. In total there were 53 (18 MC, 17 CMC, 1 CCR, and 17 OCR) science items included in PISA 2009, contained in 18 units (OECD, 2010a; OECD, 2010b).

## Reliability Estimates

Since each student was administered only one booklet they were administered different numbers of items for each subject area. The number of reading items per booklet ranged between

14 to 59 ; the number of mathematics items per booklet ranged between 11 to 27 ; and the number of science items per booklet ranged between 17 to 36 . The coefficient alpha reliability estimates for reading scores from each booklet ranged between 0.82 and 0.94 except for Booklet 12, with reliability estimates ranging from $0.73-0.75$. Reliability estimates for mathematics and science scores ranged between 0.70 and 0.90 for each booklet, except for Canadian and American mathematics scores from Booklet 9, with 0.68 and 0.65 reliabilities respectively. Most of the scores had high reliabilities. Moderate reliabilities for some of the scores were limited to only a small proportion of students ( $8 \%$ ) included in the analyses. Therefore, the inaccuracy in this study's correlational analyses due to moderate reliability of scores is expected to be minimal.

## Samples

PISA employs a two-stage stratified sampling design. In the first stage, within each jurisdiction individual schools are sampled using probability proportional to size sampling. In the second stage, 3515 -year-old students are sampled with equal probability within the sampled schools. A minimum sample size of 4,500 students in 150 schools per country was targeted by PISA. The samples for the four countries in our research ranged between 5,233 students from 165 schools in the US to 23,207 students from 978 schools in Canada (Table 1). ${ }^{1}$

In each country, 13 booklets were distributed to the examinees. One of the booklets, Booklet 6, only contained reading items, whereas all other booklets covered at least two content areas (reading + math, or reading+science, or reading + math + science). Given our focus on the relationship between reading competency and performance on either the mathematics or the

[^0]science assessment, students who responded to Booklet 6 were not included in analyses that examined correlational relationships between reading and the other two subjects. Language Group Definitions

The research focused on investigating potential threats to validity of mathematics and science score interpretations due to students' low reading proficiency levels. Therefore, the first step in our analyses was to identify groups of students with limited language proficiency levels due to their societal contexts. To identify such groups of students in the four countries included in this research we considered student responses to two variables contained in the PISA Student Questionnaire. The first variable (Question 17) asks students about their country of birth, and the second variable (Question 19) asks students what language they speak at home most of the time. We compared reading scores of four language groups that were created by using both of these variables: (a) students who were born in the country of the test and spoke English at home most of the time; (b) students who were not born in the country of the test but spoke English at home most of the time; (c) students who were born in the country of the test but spoke a different language at home most of the time; and (d) students who were not born in the country of the test and spoke a language other than English at home most of the time. A two-factor Analysis of Variance (ANOVA) (immigrant status, language at home, and immigrant status x language at home) was conducted to compare reading performances of these groups for each country. The dependent variable was a $q$ score from item response theory (IRT) based scaling from separate country analyses that ranged from -4 to +4 , with an approximate mean of 0 and standard deviation of 1 (see the score scale creation section for more details). In all four countries,
language at home was a significant factor (Australia $F(1,13804)=42.649, p<0.001$; Canada; $F(1,16831)=38.218, p<0.001 ; \operatorname{UK} F(1,11424)=57.079, p<0.001 ; \operatorname{US} F(1,5078)=31.296, p$ $<0.001$ ) with students who speak English at home most of the time scoring higher. The immigrant status was significant only in the Canadian comparison $(F(1,1)=10.357, \mathrm{p}<0.001)$ with immigrant students scoring higher. The interaction between language at home and immigrant status was significant in Australia $(F(1,13804)=7.966, p<.01)$ and in the UK $(F(1,11424)=7.121, p<.01)$. In Australia and Canada, immigrant students who speak English at home outperformed all the other three groups; in the UK and the US there were similar group difference patterns but differences were not statistically significant at the $\alpha=0.05$ level. The lowest performing group was that of immigrant students who did not speak English at home most of the time. Based on these findings, whether English was spoken at home most of the time was the key variable that distinguished students with respect to reading proficiency. A finer grouping that splits the home language groups by immigrant status, that is four groups instead of two, would be desirable. However, in such a grouping, sample sizes for some of the groups would be as low as 120 ; this, however, would prohibit conducting analyses such as differential item functioning. Therefore, we decided to focus on the home language background as the key defining variable for the language groups in all four countries resulting in two groups with students who speak English most of the time at home as English Language Background (ELB) and those who do not speak English most of the time at home as Not English Language Background (NELB). Based on the empirical evidence, home language proved to be more important than immigrant status in identifying students with limited English proficiency.

Therefore, the research focused on the differential relationships between reading proficiency and mathematics and science achievement and consistency of score meaning for ELB and NELB students.

## Differential Item Functioning Analyses

Previous research demonstrated considerable measurement incomparability between countries in international assessments (Ercikan, Roth \& Asil, in press; Kankaras \& Moores, 2013; Oliveri, Olson, Ercikan, \& Zumbo, 2012). This incomparability existed even between countries administering tests in the same language (Ercikan \& McCreith, 2002; Ercikan et al., in press; Roth et al., 2013) and between language groups within countries (Ercikan et al., in press; Kankaras \& Moores, 2013; Oliveri et al., 2012). As a first step in our analyses, we therefore conducted differential item functioning (DIF) analyses to examine comparability of items between countries and between the NELB and ELB groups within countries. It is important to identify whether item scores are comparable across groups since, if item scores are not comparable, the creation of a single scale score intended to represent all groups is not appropriate. We used a procedure developed and described by Linn and Harnisch (LH; 1981) using an IRT based approach (CTB/McGraw-Hill, 1991). The primary reason for selecting this DIF detection method was its ability to accommodate matrix sampling in PISA and utilize data across booklets. The response data from matrix-sampled assessments have large amounts of completely random missing data because students take only one of the booklets in the assessment resulting in missing data on the items that were not presented to them. Combining data across booklets results in much larger samples and therefore greater power for the statistical analyses. In addition,
this method can be used to analyze both the dichotomously-scored and polytomously-scored responses found in PISA; and it can detect both uniform DIF (equal degree of DIF across ability levels) and non-uniform DIF (unequal, or no, degree of DIF for some ability levels) (Ercikan \& McCreith, 2002). Use of other DIF detection methods is desirable to verify DIF status of items. However, the matrix sampling design in PISA creates a challenge for applying other DIF detection methods such as Mantel-Haenzsel or logistic regression.

The Linn-Harnisch DIF detection procedure computes observed and predicted mean responses for focus groups matched by the overall test score. In the IRT application of the Linn-Harnisch method, the predicted score is based on a calibration using the combined data across groups and the observed mean score is the average score for the matched ability level for the focal group. IRT parameters were calibrated using the PARDUX software
(CTB/McGraw-Hill, 1991). From the differences between the predicted and observed probabilities, a $\chi^{2}$ statistic is computed and converted to a $Z$ statistic. The DIF status of an item is determined by the statistical significance of the $Z$ statistic and an effect size based on the average difference between the predicted and observed scores, $\mathrm{p}_{\text {diff. }}$. Items with a Z statistic $<2.58$ and $\left|\mathrm{p}_{\text {diff }}\right|<0.10$ are identified as moderate DIF. Large DIF is identified by $|Z|>2.58$ and $\left|\mathrm{p}_{\text {diff }}\right|<0.10$. A negative difference implies bias against the focal group. Two sets of DIF analyses were conducted examining the appropriateness of a (a) single score scale for the four countries and (b) single score scale for NELB and ELB within countries.

## Score Scale Creation

In large-scale surveys of achievement like PISA, students take a relatively small numbers of
items in one of many booklets administered to the total sample. Plausible values are created by conditioning background variables in an effort to minimize measurement error due to small number of items. The plausible value approach used by PISA draws from a posterior distribution of $\partial$ for individuals, given that individual's item responses and background characteristics in a conditioning model (Mislevy, 1991; Monseur \& Adams, 2009). In estimation of plausible values in PISA, many background variables are included in the conditioning model to minimize measurement error. Researchers have demonstrated that inclusion of too few or too many background characteristics in the conditioning model can lead to bias in subsequent analysis, particularly when $\partial$ is an explanatory variable (Monseur \& Adams, 2009; Scofield, Junker, Taylor, \& Black, in press). In particular, the conditioning used in the estimation of plausible values may create biases in some secondary data analyses. Schofield et al. (in press) has demonstrated problems when plausible values are used as covariates, as we did in our analyses with the reading plausible values. In particular, these researchers recommend creating plausible values that use only the specific independent variables used in the secondary analysis regression model. Estimating plausible values that would not lead to biased secondary analyses is beyond the scope of this research.

Therefore, in this research we did not use the plausible values available in the PISA databases. Since students receive different booklets with different numbers and sets of items we used an IRT based scaling approach to obtain individual student $q$ scores instead of a number correct score. A simultaneous calibration procedure that combined response data across 13 booklets was used. For each country, dichotomous items were scaled using the three parameter
logistic model (3PL) (Lord, 1980) and the polytomous items were scaled using the generalized partial credit model (Muraki, 1992). The scaling analyses were conducted separately for reading, mathematics, and science. We examined item fit with the Q1 statistic (Yen, 1993) and local item dependence with Q3 statistic (Yen, 1993) to determine the appropriateness of a unidimensional model fit with the data. The results indicated satisfactory fit and unidimensionality. Separate score scales were created ranging approximately between -4 and +4 with means of 0 and standard deviation of 1 for each country. Due to high proportions of DIF items in country comparisons (see results section for details about the DIF findings), separate score scales were created for each country. However, DIF was minimal between ELB and NELB within each of the countries, therefore score scales within countries are based on a single calibration for each content area which results in scores that are comparable for NELB and ELB.

Analysis of Covariance

A key method for examining the degree to which a particular variable accounts for variation in an outcome variable is Analysis of Covariance (ANCOVA) (Maxwell, O'Callaghan, \& Delaney, 1993). This method also allows for estimating adjusted mean scores for the outcome variable when the covariate is taken into account. Reading scores served as the covariate (CV) for each of the group performance comparisons of NELB and ELB; and mathematics and science scores were the dependent variables (DV). The independent variable (IV) was a grouping variable that identified students as ELB or NELB.

Results

This research focuses on examining the relationship between reading proficiency and
performance on mathematics and science assessments and how this relationship affects comparability of scores for ELB and NELB students．The first two steps of analyses involved examining performances of ELB and NELB students on the assessments and conducting DIF analyses to determine whether single scales across countries or ELB and NELB groups within countries could be used．The findings from each step of our analyses are summarized below． Descriptive Analyses

Student responses were used to estimate their reading，mathematics and science scores． Findings summarized in Table 1 indicate significant differences between the two groups＇reading （Australia $t=4.89, p<0.001$ ；Canada：$t=4.87, p<.001$ ；UK：$t=7.07, p<.001$ ；US：$t=7.33, p<.001$ ） and science scores（Australia：$t=3.99, p<0.001$ ；Canada：$t=4.84, p<.001$ ；UK：$t=4.92, p<.001$ ；US： $t=7.12, p<.001)$ in all four countries and significant differences between the two groups＇ mathematics scores in the UK $(t=3.98, p<.001)$ and the US $(t=6.71, p<.001)$ ．Where significant differences were identified，the ELB group out－performed the NELB group． «巛巛世＂Insert Table 1 about here »＂»＂）＂»

DIF

Two sets of DIF analyses were conducted．One set examined comparability of items across the four countries in order to determine whether a single score scale can be created and used in the analyses．The second set examined the comparability of items between language groups within countries．In the DIF analyses across countries，each country was compared against the combined international group（across the four countries）where each country served as the focal group and the combined international group served as the reference group．The
findings indicated large proportions of DIF item in all three subjects (Table 2). Twenty-seven percent to $39 \%$ of the reading items, $43 \%$ to $51 \%$ of the mathematics items and $40 \%$ to $74 \%$ of the science items were identified as DIF. A great majority of these items exhibited moderate level DIF. In reading, in each country analysis, approximately half of the DIF items were in favor of the focus country. However, in mathematics, larger proportions of items were against UK and US (31\% against versus $20 \%$ in favor for UK and $26 \%$ against versus $17 \%$ in favor for US). In science, there were similar proportions of items in favor of Canada as those against it. However, whereas there were larger proportions of items in favor of Australia ( $25 \%$ in favor versus $17 \%$ against ), there were larger proportions of science DIF items against UK ( $42 \%$ against versus $32 \%$ in favor) and against US ( $23 \%$ against versus $17 \%$ in favor). These DIF results point to large degrees of measurement incomparability between countries even though in each country the assessment was administered in English. They also point to the necessity for creating separate reading, mathematics and science scales for each country.

## «巛巛«« Insert Table 2 about here »"»")"»

DIF analyses within countries between NELB and ELB groups identified a small number of items as DIF, almost all in favor of ELB (Table 3). Among the reading items, $1 \%$ were identified as DIF in favor of the ELB group in Canada and the UK, $1 \%$ of the reading items was in favor of the NELB in Canada. In mathematics, there were only $3 \%$ of the items in favor of ELB in each of the Australia and UK comparisons. In science, $3 \%$ of the items were identified as DIF in Australia, Canada and the UK in favor of ELB. Except for two items (one science item in the UK comparison and the other in the Australian comparison), all the items were identified as
moderate DIF．DIF between ELB and NELB status for items did not replicate across countries and DIF items in each of the country comparison were different items．

## «巛巛巛« Insert Table 3 about here »＂）＂）＂»＂

Given possible language effects on student performance，identifying sources of DIF between the language groups is important．However，none of the DIF items were released by OECD therefore it is not possible to review the items to investigate potential sources of DIF． Other information provided by OECD about these items provided little insights about possible sources of DIF except for one of the DIF items．The reading item that was identified as having DIF against ELB in Canada was a CR item and was therefore rated by coders．The PISA technical report（OECD，2012）indicated that there was a high degree of disagreement between coders for this item within all countries．For this reason，several countries chose not to use this item．Also consistency of the item parameters was poor for this item across countries．All items identified as DIF in the ELB and NELB comparisons were removed from scaling and score creation procedures．

ANCOVA Analyses

To determine if the CV（reading proficiency）significantly interacts with the IV（home language grouping variable），an ANCOVA model including the IV，CV and the interaction term between IV and CV was tested．An assumption required for ANCOVA analyses is the homogeneity of regression slopes．That is，the CV must not have a differential association with the DV at different levels of the IV．If this assumption does not hold，then ANCOVA results cannot be interpreted meaningfully for different levels of the IV（Henson，1998；Shadish，Cook，
\& Campbell, 2002). All the assumptions of ANCOVA including normality of residuals, homogeneity of variances, linearity of regression, independence of error terms and homogeneity of regression lines were tested. For those ANCOVA analyses where the assumption of uniformity of regression lines was violated, ANCOVA was performed separately for each level of the IV (i.e., for NELB and ELB). Since the sample sizes involved in our analyses were large, and multiple comparisons were made in this study, we adjusted the significance level to be 0.001 . This significance level was applied to all the ANCOVA analyses when assessing the statistical significance ( $p$-value).

In two out of the eight models, significant interactions between the IV and CV were identified. In the Australian analyses for science as the $\operatorname{DV}(F(342,8702)=1.42, p<0.001)$ and in the Canadian analysis for mathematics as the $\operatorname{DV}(F(409,10618)=1.25, p<0.001)$ interactions were significant, which suggested violation of the homogeneity of regression slopes assumption for ANCOVA analyses. When the homogeneity of regression slopes assumption is violated, instead of conducting the ANCOVA analysis across different groups, the relationship between the CV and the DV is examined separately for each level of the IV (Green \& Salkind, 2011). For those analyses that met the homogeneity of regression slopes assumption, the ANCOVA models were applied without the CV and IV interaction term.

The results of the ANCOVA analyses are summarized in Table 4. Reading proficiency accounts for a large proportion of variance in both mathematics (up to 43\%) and science (up to $79 \%$ ). Reading scores contribute to more variance in science than in mathematics scores. In the Canadian mathematics and Australian science analyses, the interaction between the IV and CV
was significant indicating different associations between reading and mathematics in Canada and between reading and science in Australia for the two language groups．In both of these countries， reading proficiency exhibits a stronger association with the DV for NELB（ $43 \%$ versus $39 \%$ in the Canadian and $79 \%$ versus $58 \%$ in the Australian analyses）．

## ««巛«« Insert Table 4 about here »＂»＂»＂»

## Group Differences Adjusted for Reading Proficiency

The mean scores of mathematics and science after statistically adjusting for reading scores using the ANCOVA model for each country are presented in Table 5．After adjusting scores to take reading proficiency into account using the regression lines obtained from the ANCOVA，the scores for NELB increased in both mathematics and science in every country．For ELB，on the other hand，the scores tended not to change greatly，and the changes tended to be in the opposite direction leading to a drop in the scores．Exceptions were observed in Australia where the ELB mathematics scores stayed the same and science scores showed slight increase．

The adjusted scores reflected a different pattern of group differences between ELB and NELB．In mathematics，the score gap between the groups stayed significant in favor of the ELB only in the US，whereas in Canada，the differences between the two groups became significant （even though these were not significant based on the unadjusted scores）showing higher scores for NELB．In science，the differences between NELB and ELB remained significant only in Australia，even though they were significant in all four countries based on the unadjusted scores．

The purpose of this research is to examine the relationship between reading proficiency and performance on mathematics and science assessments and how this relationship affects comparability of scores for students from ELB and those from NELB. The findings indicate that reading proficiency accounts for a large proportion of variance in both mathematics (up to 43\%) and science (up to 79\%). The results tend to be similar across countries. In mathematics, across all four countries, reading proficiency accounted for approximately $40 \%$ (ranging between $39 \%$ to $43 \%$ ). In science, reading proficiency accounted for approximately $50 \%$ of the variance (ranging between $46 \%$ to 51\%) in Canada, UK and US. However, in Australia, reading scores accounted for much larger proportions ( $58 \%$ for ELB and $79 \%$ for NELB) of science scores.. These findings point to differences in the relationship between reading proficiency and performance on science assessments in Australia, which may be due to differences in science learning and assessment in Australia compared to other countries. This and other hypotheses for such a difference will be explored in the next stage of this research. These results confirm previous research on the effects of reading proficiency on mathematics and science assessment performance (e.g., Abedi \& Gandara, 2006; Hudson, Lane \& Pullen, 2005; Noble, et al., in press) and provide an estimate of the size of possible language background effects on student performance. Statistical adjustment of scores controlling for reading proficiency indicates that scores for NELB students were indeed underestimated, a phenomenon that already has raised serious concerns (Abedi \& Gandara, 2006; Noble, et al., in press; Solano-Flores, 2008).

There is consistent evidence across the four countries that group differences in mathematics and science scores are smaller when the students' reading proficiency levels are
taken into account. In Australia, the NELB and ELB mathematics achievement differences are not statistically significant (in fact reversed, with NELB scoring higher), when reading proficiency is taken into account. In Canada, differences in science scores between these language groups disappear resulting in no achievement differences between the groups; and in mathematics, the differences are reversed, with NELB group scoring higher. In the UK, differences are no longer statistically significant between the two groups. In the US, in mathematics and in Australia in science, the statistically reliable differences between groups are smaller with ELB scoring higher when reading proficiency is taken into account. Even though the correlational relationship between reading proficiency and performance on mathematics and science assessments identified in the ANCOVA was similar for ELB and NELB in six of the eight analyses, the patterns of differences between NELB and ELB mathematics and science scores when reading proficiency is taken into account point to differences in mathematics and science score meaning for these student groups.

DIF analyses are typically accompanied by expert reviews of items for identifying sources of DIF. Such reviews result in hypotheses about sources of DIF which can be followed up with further research using think aloud protocols to gather student response process data to test these hypotheses or identify new ones (Ercikan et al., 2010). In this research, DIF analyses were conducted to examine the degree to which score scales were comparable between countries and between ELB and NELB within countries. The results indicated large degrees of DIF between countries, therefore ANCOVA analyses were conducted separately for each country. Minimal DIF was identified (at most one DIF item in each comparison, typically at moderate DIF level)
between ELB and NELB within each country.. Therefore, the kind of measurement incomparability that is identified here is not the kind that is typically identified by DIF analyses. Previous research has demonstrated differences in measurement incomparability identified at item and test levels, such as those identified by confirmatory factor analyses and DIF analysis (Ercikan \& Koh, 2005; Oliveri, Erickan \& Zumbo, 2011; Zumbo, 1999). The findings in this research point to ANCOVA as an alternative approach to examine measurement comparability at the test level.

Two limiting factors in the research may have resulted in an underestimation of potential problems with interpretation of mathematics and science scores for NELB. The first is related to reading proficiency as the only variable available to examine potential language effects. Writing proficiency in English is essential for performance on assessments with constructed response items. Inclusion of writing proficiency as an additional CV may account for greater degrees of variation in mathematics and science assessments. The second is related to low measurement accuracy for some of the scores due to small numbers of items. Low reliabilities for some of these scales are expected to have affected the correlational relationships we investigated and, in particular, may have underestimated correlational relationships, such as those between reading and each of mathematics and science scores.

An additional limiting factor is related to the self-report nature of the questionnaire data in PISA. Thus, the 15 -year-olds' reporting of whether the most commonly used language at home was English or not may include inaccuracies. There is an expectation that students' responses are influenced by how they understand the question and by their perceptions of language most
spoken at home. For example, students may not have distinguished between language spoken between their parents versus among all family members.

Finally, analysis of covariance captures correlational associations. Reading performance that accounts for variation in science and mathematics scores does not indicate that reading proficiency causes the achievement gap differences. Language group differences may be related to many socio-cultural factors including differences in socio-economic background or the quality of education that students' receive. Exploration of socio-cultural differences between ELB and NELB is one of the future directions for this research.

## Conclusions and Implications

There are major limitations in how NELB versus ELB performances are interpreted on mathematics and science assessments. Comparisons of these two groups may be made directly by comparing group performance averages for these two language groups. In direct language group comparisons, the findings from this research demonstrate that assessments may underestimate NELB's mathematics and science achievement. Often, these comparisons may be indirect for example when performances of schools, districts, states and provinces are compared. In indirect comparisons, different concentrations of NELB and ELB students in the comparison units will inevitably lead to inappropriate interpretations of group performances. For example, if effectiveness of school systems are evaluated through student achievement scores in mathematics or science-as is often the case in formal accountability models used in the US and informal accountability models used in Canada-schools or districts with high concentrations of NELB will be interpreted to be demonstrating lower achievement, therefore, poorer effectiveness
in education. Since PISA results are not used for school accountability purposes, it is important to investigate the degree to which these results would replicate in assessments that are used for accountability.

A very small number of items ( $1 \%$ to $3 \%$ ) were identified as DIF between ELB and NELB. Yet, reading scores accounted for as high as $79 \%$ of the variation in science scores, which draws attention to possible disadvantages on mathematics and science assessments for NELB students with limited reading proficiencies.. This scenario is possible if low reading proficiency disadvantages NELB across the whole mathematics and science assessments rather than on specific items due to specific vocabulary or sentence structure. To determine whether linguistic demands may have been the source of DIF we recommend a review of PISA items. The findings highlight the importance of the need to minimize the effects of language on assessments where the targeted construct for measurement is not reading. These efforts need to include thorough reviews of test items, broader reviews of cultural, language, and curricular characteristics of a test by experts and consideration of language and cultural perspectives in all stages of test development. Previous research provides guidance on how test developers can minimize such effects (Abedi, in press; Lane, in press; Solano-Flores, in press). PISA test items go through multiple phases of piloting and field-testing. However, OECD does not report any efforts to minimize language burden on the examinees. Even though the stakes for individual examinees are not high in assessments such as PISA, the inaccurate estimation of student competencies in mathematics and science is expected to impact the accuracy of overall results as well as group comparisons.

This research also highlights the need for validity evidence that the scores accurately measure the targeted constructs for NELB. For instance, cognitive interviews that focus on examinee response processes may help to identify problems with science tasks for ELL students (Noble et al., in press). Such approaches are necessary for determining linguistic and cultural aspects of items that may contribute to measurement error for NELB examinees.

Finally, with respect to how scores are reported, the limitations of interpreting scores for NELB should be addressed.. We recommend three approaches. One approach may be to include measurement error due to language group membership as part of the overall measurement error. A probabilistic approach to assessment recognizes uncertainty as a result of the multiple linguistic factors that shape the ways in which students make sense of items (Solano-Flores, in press). Based on this approach, language can be considered as a source of measurement error. Estimating this type of measurement error in G-studies (see Solano-Flores, 2009) and including error as part of the score reports is likely to lead to more cautious interpretation of scores.

A second strategy involves indicating NELB status of students and providing cautionary statements about interpretation of scores in score reports. Such statements should caution users of test scores that NELB students' mathematics and science may be underestimated and that they should not be compared with scores of ELB students.

A third strategy we recommend is an extension of the second strategy and involves measuring and reporting of language proficiencies of NELB along with mathematics and science achievement scores. Recommendations can be made about at what language proficiency level scores can be meaningfully interpreted. Such recommendations will need to be based on
empirical evidence to determine a language proficiency level beyond which mathematics and science scores are underestimated.

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Table 1
Reading and Mathematics Mean Scores and t-test Results

| Country | Subject | ELB |  |  | NELB |  |  | t-test |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\mathrm{N}_{1}$ | Mean | S. D. | $\mathrm{N}_{2}$ | Mean | S. D. | $t$ | $p$ |
| Australia | Reading | 12756 | 0.03 | 1.14 | 1124 | -0.17 | 1.33 | 4.89 | $<0.001$ |
|  | Mathematics | 8852 | 0.02 | 0.98 | 786 | 0.01 | 1.10 | 0.58 | 0.561 |
|  | Science | 8818 | 0.06 | 1.20 | 802 | -0.16 | 1.47 | 3.99 | <0.001 |
| Canada | Reading | 15091 | 0.05 | 1.14 | 1745 | -0.11 | 1.28 | 4.87 | $<0.001$ |
|  | Mathematics | 10437 | -0.05 | 1.32 | 1190 | -0.05 | 1.45 | 0.18 | 0.858 |
|  | Science | 10454 | 0.04 | 0.97 | 1192 | 09.11 | 1.08 | 4.84 | <0.001 |
| UK | Reading | 11018 | 0.01 | 1.18 | 526 | -0.42 | 1.40 | 7.07 | $<0.001$ |
|  | Mathematics | 7615 | 0.03 | 0.99 | 373 | -0.18 | 1.03 | 3.98 | <0.001 |
|  | Science | 7612 | 0.05 | 1.20 | 365 | -0.31 | 1.37 | 4.92 | <0.001 |
| US | Reading | 4466 | 0.03 | 1.22 | 669 | -0.34 | 1.17 | 7.33 | $<0.001$ |
|  | Mathematics | 3097 | 0.05 | 1.00 | 481 | -0.28 | 0.92 | 6.71 | <0.001 |
|  | Science | 3095 | 0.00 | 1.27 | 459 | -0.37 | 1.25 | 7.12 | <0.001 |

Note: ELB= English language background, NELB= Non-English Language Background

Table 2
Number and percentage of DIF items between countries by subject

| Country | Direction | Reading <br> $(177 ~ i t e m s) ~$ |  | Mathematics <br> (35 items) |  | Science <br> (53 items) |  |
| :--- | :--- | :--- | :---: | :---: | :---: | :---: | :---: |
|  |  | Moderate | Large | Moderate | Large | Moderate | Large |
| Australia | In favour | $22(19 \%)$ | 0 | $8(23 \%)$ | 0 | $13(25 \%)$ | 0 |
|  | Against | $21(18 \%)$ | 0 | $8(23 \%)$ | 0 | $9(17 \%)$ | 0 |
| Canada | In favour | $14(12 \%)$ | 0 | $8(23 \%)$ | 0 | $15(28 \%)$ | 0 |
|  | Against | $17(15 \%)$ | 0 | $8(23 \%)$ | 0 | $16(30 \%)$ | 0 |
| UK | In favour | $22(19 \%)$ | 0 | $7(20 \%)$ | 0 | $14(26 \%)$ | $3(6 \%)$ |
|  | Against | $21(18 \%)$ | $2(2 \%)$ | $11(31 \%)$ | 0 | $21(40 \%)$ | $1(2 \%)$ |
| US | In favour | $16(14 \%)$ | $1(1 \%)$ | $6(17 \%)$ | 0 | $9(17 \%)$ | 0 |
|  | Against | $18(16 \%)$ | 0 | $8(23 \%)$ | $1(3 \%)$ | $12(23 \%)$ | 0 |

Table 3
Number and percentage of DIF items between ELB and NELB groups within countries

| Country | Favouring | Reading <br> (166 items) |  | Mathematics <br> (35 items) |  | Science <br> (53 items) |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  | Moderate | Large | Moderate | Large | Moderate | Large |
| Australia | ELB | 0 | 0 | $1(3 \%)$ | 0 | 0 | $1(3 \%)$ |
|  | NELB | 0 | 0 | 0 | 0 | 0 | 0 |
| Canada | ELB | $1(1 \%)$ | 0 | 0 | 0 | $1(3 \%)$ | 0 |
|  | NELB | $1(1 \%)$ | 0 | 0 | 0 | 0 | 0 |
| UK | ELB | $1(1 \%)$ | 0 | $1(3 \%)$ | 0 | 0 | $1(3 \%)$ |
|  | NELB | 0 | 0 | 0 | 0 | 0 | 0 |
| US | ELB | 0 | 0 | 0 | 0 | 0 | 0 |
|  | NELB | 0 | 0 | 0 | 0 | 0 | 0 |

Table 4
ANCOVA Results

| DV | Country | Variable | F | $p$ | Partial Eta <br> Square |
| :--- | :--- | :--- | ---: | :---: | :---: |
| Mathematics | Australia | Home Language | 9.83 | 0.002 | 0.001 |
|  |  | Reading | 7028.25 | $<0.001$ | 0.422 |
|  | Canada | Reading (ELB) | $\mathbf{6 6 4 9 . 0 7}$ | $<\mathbf{0 . 0 0 1}$ | $\mathbf{0 . 3 8 9}$ |
|  |  | Reading (NELB) | $\mathbf{8 8 8 . 7 9}$ | $<\mathbf{0 . 0 0 1}$ | $\mathbf{0 . 4 2 8}$ |
|  | UK | Home Language | 0.01 | 0.913 | 0.000 |
|  |  | Reading | 5106.21 | $<0.001$ | 0.390 |
|  |  | Home Language | 15.78 | $<0.001$ | 0.004 |
| Science | Australia | Reading (ELB) | 2271.38 | $<0.001$ | 0.389 |
|  |  | Reading (NELB) | $\mathbf{2 0 . 2 1}$ | $<\mathbf{0 . 0 0 1}$ | $\mathbf{0 . 5 8 0}$ |
|  |  | Reading | $\mathbf{4 . 9 1}$ | $<\mathbf{0 . 0 0 1}$ | $\mathbf{0 . 7 9 4}$ |
|  | Canada | Home Language | 7.85 | 0.005 | 0.001 |
|  |  | Reading | 10098.27 | $<0.001$ | 0.464 |
|  | UK | Home Language | 1.26 | 0.213 | 0.000 |
|  |  | Reading | 8284.00 | $<0.001$ | 0.510 |
|  | US | Home Language | 18.60 | $<0.001$ | 0.005 |
|  |  | Reading | 3125.45 | $<0.001$ | 0.468 |

Note: ANCOVA results highlighted in bold were estimated within each group separately since the assumption of uniformity of regression slopes was not met. This is also the reason for why separate fit statistic, significance level and effect size are reported for ELB and NELB in Canada (Mathematics) and Australia (Science)

Table 5
Adjusted and Unadjusted Means for each Group

| Country | Language <br> Group | Mathematics |  | Science |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Adjusted <br> Mean (SE) | Unadjusted <br> Mean (SE) | Adjusted <br> Mean (SE) | Unadjusted <br> Mean (SE) |
| Australia | ELB | $\begin{gathered} 0.02 \\ (0.008) \end{gathered}$ | $\begin{gathered} 0.02 \\ (0.010) \end{gathered}$ | $\begin{gathered} 0.10^{*} \\ (0.009) \end{gathered}$ | $\begin{gathered} 0.06^{*} \\ (0.013) \end{gathered}$ |
|  | NELB | $\begin{gathered} 0.10 \\ (0.027) \end{gathered}$ | $\begin{gathered} 0.01 \\ (0.039) \end{gathered}$ | $\begin{gathered} -0.12 * \\ (0.039) \end{gathered}$ | $\begin{aligned} & -0.16^{*} \\ & (0.052) \end{aligned}$ |
| Canada | ELB | $\begin{gathered} -0.13^{*} \\ (0.008) \\ \hline \end{gathered}$ | $\begin{gathered} -0.05 \\ (0.013) \end{gathered}$ | $\begin{gathered} 0.03 \\ (0.007) \end{gathered}$ | $\begin{gathered} 0.04^{*} \\ (0.010) \end{gathered}$ |
|  | NELB | $\begin{gathered} -0.03^{*} \\ (0.028) \end{gathered}$ | $\begin{gathered} -0.05 \\ (0.042) \end{gathered}$ | $\begin{gathered} -0.03 \\ (0.021) \end{gathered}$ | $\begin{gathered} -0.11^{*} \\ (0.031) \end{gathered}$ |
| UK | ELB | $\begin{gathered} 0.02 \\ (0.009) \end{gathered}$ | $\begin{gathered} 0.03^{*} \\ (0.011) \end{gathered}$ | $\begin{gathered} 0.04 \\ (0.010) \end{gathered}$ | $\begin{gathered} 0.05^{*} \\ (0.014) \end{gathered}$ |
|  | NELB | $\begin{gathered} 0.01 \\ (0.040) \end{gathered}$ | $\begin{gathered} -0.18^{*} \\ (0.053) \end{gathered}$ | $\begin{gathered} -0.02 \\ (0.044) \end{gathered}$ | $\begin{aligned} & -0.31^{*} \\ & (0.071) \end{aligned}$ |
| US | ELB | $\begin{gathered} 0.03 * \\ (0.014) \end{gathered}$ | $\begin{gathered} 0.05^{*} \\ (0.018) \end{gathered}$ | $\begin{gathered} 0.03 \\ (0.013) \end{gathered}$ | $\begin{gathered} 0.06^{*} \\ (0.018) \end{gathered}$ |
|  | NELB | $\begin{aligned} & -0.13^{*} \\ & (0.036) \end{aligned}$ | $\begin{aligned} & -0.28^{*} \\ & (0.042) \end{aligned}$ | $\begin{gathered} -0.13 \\ (0.034) \end{gathered}$ | $\begin{aligned} & -0.30^{*} \\ & (0.045) \end{aligned}$ |


[^0]:    ${ }^{1}$ Only the students who took PISA in English in the four countries were included in our study.

